

Random Vectors

• In principle, going from a pair of random variables X and Y to n random variables X_1, X_2, \dots, X_n is easy.

→ The joint cumulative distribution function (CDF) is

$$F_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n) = \mathbb{P}[X_1 \leq x_1, X_2 \leq x_2, \dots, X_n \leq x_n]$$

→ For discrete random variables, the joint probability mass function (PMF) is

$$P_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n) = \mathbb{P}[X_1 = x_1, X_2 = x_2, \dots, X_n = x_n]$$

→ For continuous random variables, we have a joint probability density function (PDF) satisfying

$$F_{X_1, X_2, \dots, X_n}(x_1, x_2, \dots, x_n) = \int_{-\infty}^{x_n} \dots \int_{-\infty}^{x_2} \int_{-\infty}^{x_1} f_{X_1, X_2, \dots, X_n}(u_1, u_2, \dots, u_n) du_1 du_2 \dots du_n$$

• The usual PMF/PDF properties hold:

→ Non-negativity: $P_{x_1, \dots, x_n}(x_1, \dots, x_n) \geq 0$

$$f_{x_1, \dots, x_n}(x_1, \dots, x_n) \geq 0$$

→ Normalization: $\sum_{x_1 \in R_{x_1}} \dots \sum_{x_n \in R_{x_n}} P_{x_1, \dots, x_n}(x_1, \dots, x_n) = 1$

$$\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f_{x_1, \dots, x_n}(x_1, \dots, x_n) dx_1 \dots dx_n = 1$$

→ Additivity: $\mathbb{P}[\{(x_1, \dots, x_n) \in B\}] = \sum_{(x_1, \dots, x_n) \in B} P_{x_1, \dots, x_n}(x_1, \dots, x_n)$

$$\mathbb{P}[\{(x_1, \dots, x_n) \in B\}] = \int_B \dots \int f_{x_1, \dots, x_n}(x_1, \dots, x_n) dx_1 \dots dx_n$$

• x_1, \dots, x_n are independent if $P_{x_1, \dots, x_n}(x_1, \dots, x_n) = P_{x_1}(x_1) \dots P_{x_n}(x_n)$

$$f_{x_1, \dots, x_n}(x_1, \dots, x_n) = f_{x_1}(x_1) \dots f_{x_n}(x_n)$$

- The expected value $\mathbb{E}[g(x_1, \dots, x_n)]$ of a function $g(x_1, \dots, x_n)$ of n random variables is

$$\mathbb{E}[g(x_1, \dots, x_n)] = \sum_{x_1 \in R_{x_1}} \cdots \sum_{x_n \in R_{x_n}} g(x_1, \dots, x_n) P_{x_1, \dots, x_n}(x_1, \dots, x_n)$$

(Discrete)

$$\mathbb{E}[g(x_1, \dots, x_n)] = \int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} g(x_1, \dots, x_n) f_{x_1, \dots, x_n}(x_1, \dots, x_n) dx_1 \cdots dx_n$$

(Continuous)

- Linearity of Expectation:** For any functions $g_1(x_1, \dots, x_n), \dots, g_m(x_1, \dots, x_n)$ and constants a_1, \dots, a_m ,

$$\begin{aligned} & \mathbb{E}[a_1 g_1(x_1, \dots, x_n) + \cdots + a_m g_m(x_1, \dots, x_n)] \\ &= a_1 \mathbb{E}[g_1(x_1, \dots, x_n)] + \cdots + a_m \mathbb{E}[g_m(x_1, \dots, x_n)] \end{aligned}$$

• **Conditioning** also generalizes naturally. Say we want the conditional distribution of X_1, \dots, X_m given X_{m+1}, \dots, X_n . We just divide the joint distribution by the marginal distribution:

→ The **conditional PMF** of X_1, \dots, X_m given X_{m+1}, \dots, X_n is

$$P_{X_1, \dots, X_m | X_{m+1}, \dots, X_n}(x_1, \dots, x_m | x_{m+1}, \dots, x_n) = \begin{cases} \frac{P_{X_1, \dots, X_n}(x_1, \dots, x_n)}{P_{X_{m+1}, \dots, X_n}(x_{m+1}, \dots, x_n)} & \text{for } (x_1, \dots, x_n) \\ & \text{in } R_{X_1, \dots, X_n} \\ 0 & \text{otherwise} \end{cases}$$

→ The **conditional PDF** of X_1, \dots, X_m given X_{m+1}, \dots, X_n is

$$f_{X_1, \dots, X_m | X_{m+1}, \dots, X_n}(x_1, \dots, x_m | x_{m+1}, \dots, x_n) = \begin{cases} \frac{f_{X_1, \dots, X_n}(x_1, \dots, x_n)}{f_{X_{m+1}, \dots, X_n}(x_{m+1}, \dots, x_n)} & \text{for } (x_1, \dots, x_n) \\ & \text{in } R_{X_1, \dots, X_n} \\ 0 & \text{otherwise} \end{cases}$$

→ The **conditional expected value** $E[g(X_1, \dots, X_n) | X_{m+1} = x_{m+1}, \dots, X_n = x_n]$ is

Discrete: $\sum_{x_1 \in R_{X_1}} \dots \sum_{x_m \in R_{X_m}} g(x_1, \dots, x_m, x_{m+1}, \dots, x_n) P_{X_1, \dots, X_m | X_{m+1}, \dots, X_n}(x_1, \dots, x_m | x_{m+1}, \dots, x_n)$

Continuous: $\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} g(x_1, \dots, x_m, x_{m+1}, \dots, x_n) f_{X_1, \dots, X_m | X_{m+1}, \dots, X_n}(x_1, \dots, x_m | x_{m+1}, \dots, x_n) dx_1 \dots dx_m$

- The main issue with $n > 2$ random variables is that **working out n -dimensional sums and integrals is hard.**
- However, in some cases, we can get by with only working out first- and second-order statistics, such as the mean and covariance.
- For n random variables X_1, \dots, X_n , we would need
 - n means $\mathbb{E}[X_1], \dots, \mathbb{E}[X_n]$
 - n^2 covariances $\text{Cov}[X_i, X_j] = \mathbb{E}[(X_i - \mathbb{E}[X_i])(X_j - \mathbb{E}[X_j])]$
for $i, j \in \{1, \dots, n\}$
(Note that $\text{Cov}[X_i, X_i] = \text{Var}[X_i]$.)

These are just one- and two-dimensional calculations (and relatively easy to learn from data).

- Why are the means and covariances useful?
 - One important reason: Easy to work out how means and covariances evolve under **linear transformations**.
 - Another reason: Jointly Gaussian random variables are fully specified by their means and covariances.
- To explore these ideas, we need vectors and matrices.
- A **random vector** \underline{X} is a column vector whose entries are random variables.

→ Notation: $\underline{X} = \begin{bmatrix} X_1 \\ \vdots \\ X_n \end{bmatrix}$ **random vector**

$\underline{x} = \begin{bmatrix} x_1 \\ \vdots \\ x_n \end{bmatrix}$ **value \underline{X} takes**

$P_{\underline{X}}(\underline{x}) = P_{X_1, \dots, X_n}(x_1, \dots, x_n)$ **Joint PMF**

$f_{\underline{X}}(\underline{x}) = f_{X_1, \dots, X_n}(x_1, \dots, x_n)$ **Joint PDF**

- We can organize the means and covariances into vectors and matrices.
- The **mean vector** $\underline{\mu}_x$ of a random vector \underline{X} is the column vector whose entries are the expected values of the corresponding entries of \underline{X} :

$$\underline{\mu}_x = \mathbb{E}[\underline{X}] = \begin{bmatrix} \mathbb{E}[X_1] \\ \vdots \\ \mathbb{E}[X_n] \end{bmatrix}$$

In other words, to take the expectation of a random vector, just take the expectation of each entry.

- **Linearity of Expectation:** For any (constant) matrix \mathbf{A} and vector \underline{b} ,

$$\mathbb{E}[\mathbf{A}\underline{X} + \underline{b}] = \mathbf{A}\mathbb{E}[\underline{X}] + \underline{b}$$

- The covariance matrix $\Sigma_{\underline{x}}$ of a random vector \underline{x} is

$$\Sigma_{\underline{x}} = \mathbb{E}[(\underline{x} - \mathbb{E}[\underline{x}])(\underline{x} - \mathbb{E}[\underline{x}])^T]$$

$$= \mathbb{E} \left[\begin{bmatrix} X_1 - \mathbb{E}[X_1] \\ X_2 - \mathbb{E}[X_2] \\ \vdots \\ X_n - \mathbb{E}[X_n] \end{bmatrix} \begin{bmatrix} X_1 - \mathbb{E}[X_1] & X_2 - \mathbb{E}[X_2] & \dots & X_n - \mathbb{E}[X_n] \end{bmatrix} \right]$$

$(i, j)^{\text{th}}$ entry contains $\text{Cov}[X_i, X_j]$

$$= \begin{bmatrix} \text{Var}[X_1] & \text{Cov}[X_1, X_2] & \dots & \text{Cov}[X_1, X_n] \\ \text{Cov}[X_2, X_1] & \text{Var}[X_2] & \dots & \text{Cov}[X_2, X_n] \\ \vdots & \vdots & \ddots & \vdots \\ \text{Cov}[X_n, X_1] & \text{Cov}[X_n, X_2] & \dots & \text{Var}[X_n] \end{bmatrix}$$

- Alternate Covariance Matrix Equation:

$$\Sigma_{\underline{x}} = \mathbb{E}[\underline{x}\underline{x}^T] - \mathbb{E}[\underline{x}](\mathbb{E}[\underline{x}])^T$$

- Covariance Matrix Properties: The covariance matrix $\Sigma_{\underline{x}}$
 - is **symmetric** $\Sigma_{\underline{x}} = \Sigma_{\underline{x}}^T$ (since $\text{Cov}[x_i, x_j] = \text{Cov}[x_j, x_i]$)
 - is **positive semi-definite**, $\underline{a}^T \Sigma_{\underline{x}} \underline{a} \geq 0$ for any vector \underline{a}
 - has all **real, non-negative eigenvalues**
 - has **n distinct eigenvectors**, each **perpendicular** to the others

These properties will be useful later in the course.

• Covariance Matrix after a Linear Transformation:

Let \underline{X} be a random vector with covariance matrix $\Sigma_{\underline{X}}$ and let $\underline{Y} = \mathbf{A} \underline{X} + \underline{b}$. Then, the covariance matrix of \underline{Y} is $\Sigma_{\underline{Y}} = \mathbf{A} \Sigma_{\underline{X}} \mathbf{A}^T$.

Why?

$$\Sigma_{\underline{Y}} = \mathbb{E}[(\underline{Y} - \mathbb{E}[\underline{Y}])(\underline{Y} - \mathbb{E}[\underline{Y}])^T]$$

Linearity of Expectation

$$= \mathbb{E}[(\mathbf{A} \underline{X} + \underline{b} - \mathbb{E}[\mathbf{A} \underline{X} + \underline{b}])(\mathbf{A} \underline{X} + \underline{b} - \mathbb{E}[\mathbf{A} \underline{X} + \underline{b}])^T]$$

$$= \mathbb{E}[(\mathbf{A} \underline{X} + \underline{b} - \mathbf{A} \mathbb{E}[\underline{X}] - \underline{b})(\mathbf{A} \underline{X} + \underline{b} - \mathbf{A} \mathbb{E}[\underline{X}] - \underline{b})^T]$$

Linearity of Expectation (twice)

$$= \mathbb{E}[\mathbf{A}(\underline{X} - \mathbb{E}[\underline{X}])(\underline{X} - \mathbb{E}[\underline{X}])^T \mathbf{A}^T] \quad \leftarrow (\mathbf{AC})^T = \mathbf{C}^T \mathbf{A}^T$$

$$= \mathbf{A} \mathbb{E}[(\underline{X} - \mathbb{E}[\underline{X}])(\underline{X} - \mathbb{E}[\underline{X}])^T] \mathbf{A}^T$$

$$= \mathbf{A} \Sigma_{\underline{X}} \mathbf{A}^T$$

• Example: \underline{X} is a random vector with mean vector $\underline{\mu}_X = \begin{bmatrix} 1 \\ -1 \end{bmatrix}$ and covariance matrix $\underline{\Sigma}_X = \begin{bmatrix} 2 & 1 \\ 1 & 2 \end{bmatrix}$.

Let $\underline{Y} = \begin{bmatrix} -1 & 1 \\ 3 & 2 \end{bmatrix} \underline{X} + \begin{bmatrix} 4 \\ -3 \end{bmatrix}$. Determine $\underline{\mu}_Y$ and $\underline{\Sigma}_Y$.

Linearity of Expectation: $\mathbb{E}[\underline{A}\underline{X} + \underline{b}] = \underline{A}\mathbb{E}[\underline{X}] + \underline{b}$

$$\underline{\mu}_Y = \begin{bmatrix} -1 & 1 \\ 3 & 2 \end{bmatrix} \begin{bmatrix} 1 \\ -1 \end{bmatrix} + \begin{bmatrix} 4 \\ -3 \end{bmatrix} = \begin{bmatrix} -1 \cdot 1 + 1 \cdot (-1) \\ 3 \cdot 1 + 2 \cdot (-1) \end{bmatrix} + \begin{bmatrix} 4 \\ -3 \end{bmatrix} = \begin{bmatrix} -2 \\ 1 \end{bmatrix} + \begin{bmatrix} 4 \\ -3 \end{bmatrix} = \begin{bmatrix} 2 \\ -2 \end{bmatrix}$$

Covariance of a Linear Transformation: $\underline{\Sigma}_Y = \underline{A}\underline{\Sigma}_X\underline{A}^T$

$$\begin{aligned} \underline{\Sigma}_Y &= \begin{bmatrix} -1 & 1 \\ 3 & 2 \end{bmatrix} \begin{bmatrix} 2 & 1 \\ 1 & 2 \end{bmatrix} \begin{bmatrix} -1 & 3 \\ 1 & 2 \end{bmatrix} = \begin{bmatrix} -1 \cdot 2 + 1 \cdot 1 & -1 \cdot 1 + 1 \cdot 2 \\ 3 \cdot 2 + 2 \cdot 1 & 3 \cdot 1 + 2 \cdot 2 \end{bmatrix} \begin{bmatrix} -1 & 3 \\ 1 & 2 \end{bmatrix} \\ &= \begin{bmatrix} -1 & 1 \\ 8 & 7 \end{bmatrix} \begin{bmatrix} -1 & 3 \\ 1 & 2 \end{bmatrix} = \begin{bmatrix} -1 \cdot (-1) + 1 \cdot 1 & -1 \cdot 3 + 1 \cdot 2 \\ 8 \cdot (-1) + 7 \cdot 1 & 8 \cdot 3 + 7 \cdot 2 \end{bmatrix} = \begin{bmatrix} 1 & -1 \\ -1 & 38 \end{bmatrix} \end{aligned}$$