

## Likelihood Ratio

- Recall the framework of binary hypothesis testing:
  - Two hypotheses  $H_0$  and  $H_1$  that partition  $\Omega$ .
  - An observation  $Y$  whose values are drawn according to

### Discrete Case

$P_{Y|H_0}(y)$  if  $H_0$  occurs

$P_{Y|H_1}(y)$  if  $H_1$  occurs

### Continuous Case

$f_{Y|H_0}(y)$  if  $H_0$  occurs

$f_{Y|H_1}(y)$  if  $H_1$  occurs

→ A decision rule  $D(Y)$  that outputs 0 or 1 as its guess for the hypothesis based only on  $Y$ .

- We will revisit the ML and MAP rules from the perspective of the likelihood ratio  $L(y)$

$$L(y) = \begin{cases} \frac{P_{Y|H_1}(y)}{P_{Y|H_0}(y)} & Y \text{ is discrete} \\ \frac{f_{Y|H_1}(y)}{f_{Y|H_0}(y)} & Y \text{ is continuous} \end{cases}$$

- Maximum Likelihood (ML) Rule:

$$D^{ML}(y) = \begin{cases} 1, & L(y) \geq 1 \\ 0, & L(y) < 1 \end{cases} = \begin{cases} 1, & \ln(L(y)) \geq 0 \\ 0, & \ln(L(y)) < 0 \end{cases}$$

→  $\ln(L(y))$  is called the **log-likelihood ratio** and can sometimes help simplify the form of the decision rule.

- Maximum a Posteriori (MAP) Rule:

$$D^{MAP}(y) = \begin{cases} 1, & L(y) \geq \frac{P[H_0]}{P[H_1]} \\ 0, & L(y) < \frac{P[H_0]}{P[H_1]} \end{cases} = \begin{cases} 1, & \ln(L(y)) \geq \ln\left(\frac{P[H_0]}{P[H_1]}\right) \\ 0, & \ln(L(y)) < \ln\left(\frac{P[H_0]}{P[H_1]}\right) \end{cases}$$

→ Why?  $P_{Y|H_1}(y) P[H_1] \geq P_{Y|H_0}(y) P[H_0] \Leftrightarrow \frac{P_{Y|H_1}(y)}{P_{Y|H_0}(y)} \geq \frac{P[H_0]}{P[H_1]}$

$L(y) \rightarrow$

Discrete Case

$$L(y) = \frac{P_{Y|H_1}(y)}{P_{Y|H_0}(y)}$$

Continuous Case

$$L(y) = \frac{f_{Y|H_1}(y)}{f_{Y|H_0}(y)}$$

• Example: Given  $H_0$  occurs,  $Y$  is Binomial  $(n, p_0)$ .

Given  $H_1$  occurs,  $Y$  is Binomial  $(n, p_1)$ .

$$L(y) = \frac{P_{Y|H_1}(y)}{P_{Y|H_0}(y)} = \frac{\binom{n}{y} p_1^y (1-p_1)^{n-y}}{\binom{n}{y} p_0^y (1-p_0)^{n-y}} = \left( \frac{p_1(1-p_0)}{p_0(1-p_1)} \right)^y \frac{(1-p_1)^n}{(1-p_0)^n}$$

$$D^{ML}(y) = \begin{cases} 1, & L(y) \geq 1 \\ 0, & L(y) < 1 \end{cases} = \begin{cases} 1, & \left( \frac{p_1(1-p_0)}{p_0(1-p_1)} \right)^y \frac{(1-p_1)^n}{(1-p_0)^n} \geq 1 \\ 0, & \left( \frac{p_1(1-p_0)}{p_0(1-p_1)} \right)^y \frac{(1-p_1)^n}{(1-p_0)^n} < 1 \end{cases}$$

Can simplify using the log-likelihood ratio.

$$\ln(L(y)) = \ln \left( \left( \frac{p_1(1-p_0)}{p_0(1-p_1)} \right)^y \frac{(1-p_1)^n}{(1-p_0)^n} \right) = y \ln \left( \frac{p_1(1-p_0)}{p_0(1-p_1)} \right) + n \ln \left( \frac{1-p_1}{1-p_0} \right)$$

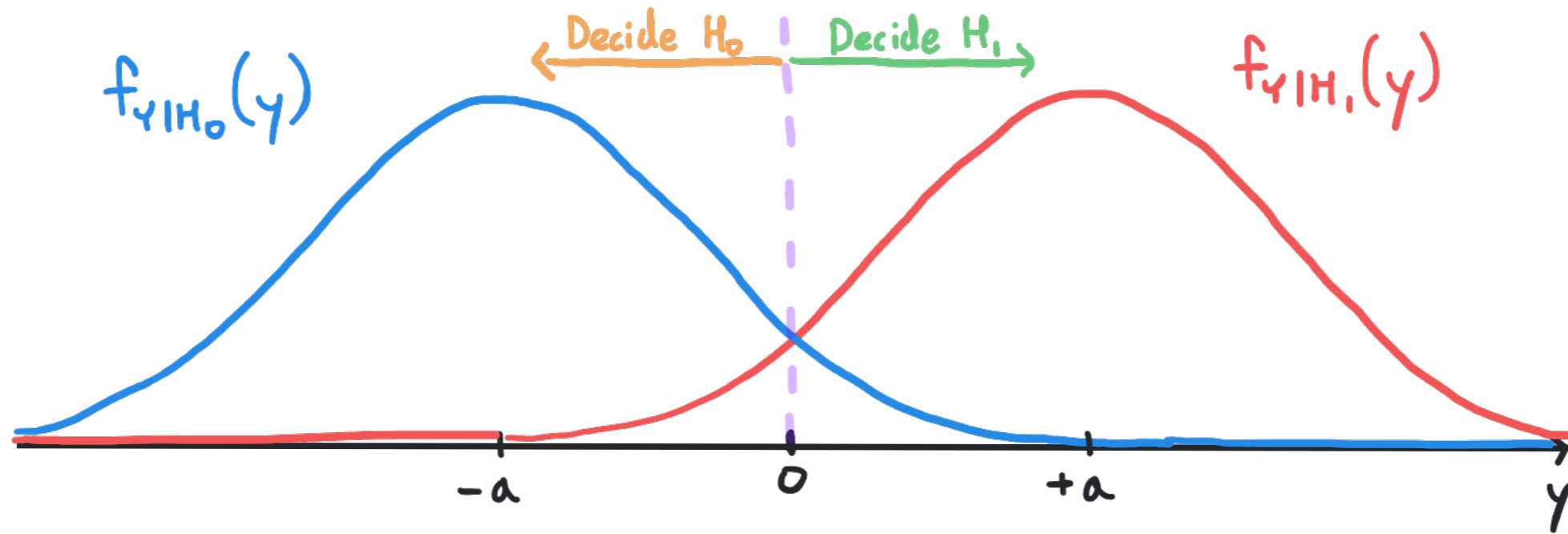
$$D^{ML}(y) = \begin{cases} 1, & \ln(L(y)) \geq 0 \\ 0, & \ln(L(y)) < 0 \end{cases} = \begin{cases} 1, & y \ln \left( \frac{p_1(1-p_0)}{p_0(1-p_1)} \right) \geq -n \ln \left( \frac{1-p_1}{1-p_0} \right) \\ 0, & y \ln \left( \frac{p_1(1-p_0)}{p_0(1-p_1)} \right) < -n \ln \left( \frac{1-p_1}{1-p_0} \right) \end{cases}$$

In the previous video,  $n=3$ ,  $p_0 = \frac{1}{2}$ ,  $p_1 = \frac{3}{4}$ .

$$y \ln \left( \frac{\frac{3}{4}(1-\frac{1}{2})}{\frac{1}{2}(1-\frac{3}{4})} \right) \geq -3 \ln \left( \frac{1-\frac{3}{4}}{1-\frac{1}{2}} \right) \Rightarrow y \ln(3) \geq -3 \ln\left(\frac{1}{2}\right) \Rightarrow y \geq 1.89 \Rightarrow y \in \{2, 3\}$$

Same decision rule!

- Example: Given  $H_0$  occurs,  $Y$  is Gaussian  $(-a, \sigma^2)$ . Assume  $a > 0$ .  
Given  $H_1$  occurs,  $Y$  is Gaussian  $(+a, \sigma^2)$ .

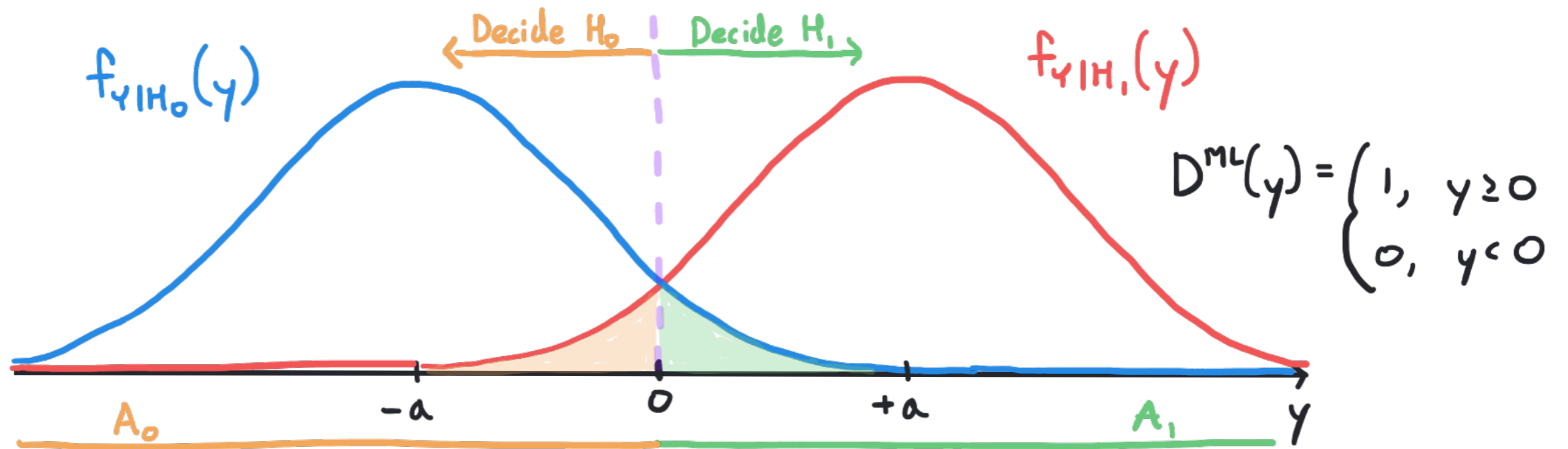


→ From the plot, we can guess that the ML rule decides  $H_1$  when  $y \geq 0$  and  $H_0$  when  $y < 0$ . Check using log-likelihood ratio.

$$\begin{aligned} \ln(L(y)) &= \ln\left(\frac{\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y-a)^2}{2\sigma^2}\right)}{\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y+a)^2}{2\sigma^2}\right)}\right) \\ &= \ln\left(\exp\left(-\frac{1}{2\sigma^2}(y^2 - 2ay + a^2) + \frac{1}{2\sigma^2}(y^2 + 2ay + a^2)\right)\right) = \frac{2ay}{\sigma^2} \end{aligned}$$

$$D^{ML}(y) = \begin{cases} 1, & \ln(L(y)) \geq 0 \\ 0, & \ln(L(y)) < 0 \end{cases} = \begin{cases} 1, & \frac{2ay}{\sigma^2} \geq 0 \\ 0, & \frac{2ay}{\sigma^2} < 0 \end{cases} = \begin{cases} 1, & y \geq 0 \\ 0, & y < 0 \end{cases} \quad \checkmark$$

- Example: Given  $H_0$  occurs,  $Y$  is Gaussian  $(-a, \sigma^2)$ . Assume  $a > 0$ .  
Given  $H_1$  occurs,  $Y$  is Gaussian  $(+a, \sigma^2)$ .



→ What is the probability of error for the ML rule?

$$P_e = P_{FA} IP[H_0] + P_{MD} IP[H_1]$$

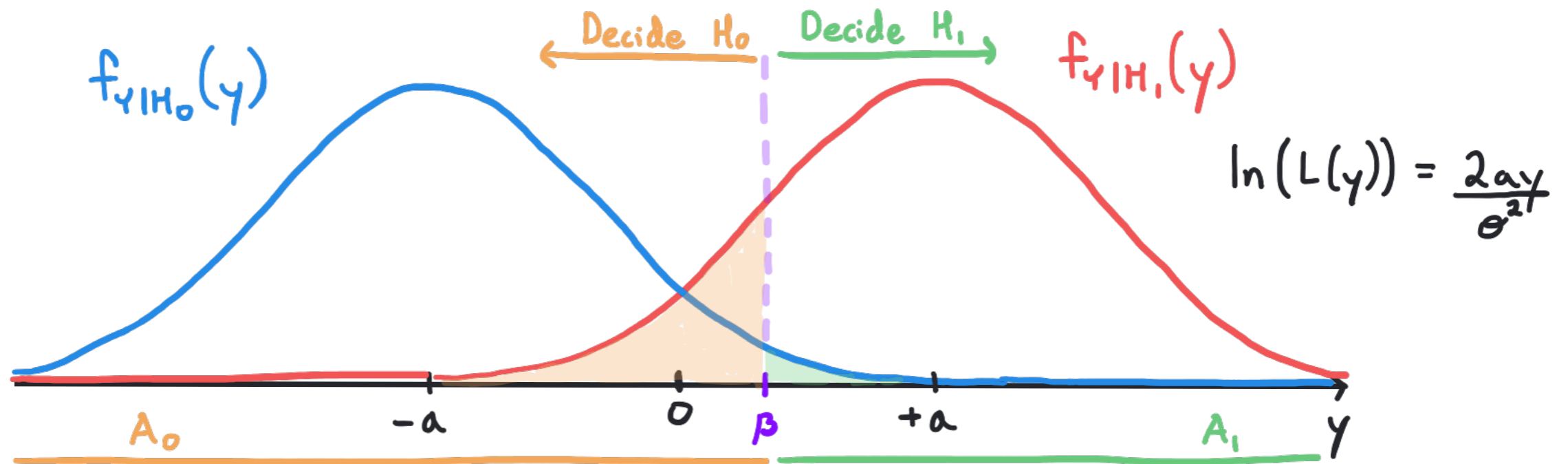
$$P_{FA} = IP[\{Y \in A_1\} | H_0] = IP[Y \geq 0 | H_0] = \int_0^{\infty} f_{Y|H_0}(y) dy = Q\left(\frac{0 - (-a)}{\sigma}\right) = Q\left(\frac{a}{\sigma}\right)$$

$$P_{MD} = IP[\{Y \in A_0\} | H_1] = IP[Y < 0 | H_1] = \int_{-\infty}^0 f_{Y|H_1}(y) dy = \Phi\left(\frac{0 - (a)}{\sigma}\right) = Q\left(\frac{a}{\sigma}\right)$$

$$P_e = Q\left(\frac{a}{\sigma}\right) IP[H_0] + Q\left(\frac{a}{\sigma}\right) IP[H_1] = Q\left(\frac{a}{\sigma}\right)$$

In this special case, we do not need  $IP[H_0]$ ,  $IP[H_1]$  explicitly.

- Example: Given  $H_0$  occurs,  $Y$  is Gaussian  $(-a, \sigma^2)$ . Assume  $a > 0$ .  
Given  $H_1$  occurs,  $Y$  is Gaussian  $(+a, \sigma^2)$ .



→ What is the MAP rule and its probability of error?

$$D^{\text{MAP}}(y) = \begin{cases} 1, & \ln(L(y)) \geq \ln\left(\frac{P[H_0]}{P[H_1]}\right) \\ 0, & \ln(L(y)) < \ln\left(\frac{P[H_0]}{P[H_1]}\right) \end{cases} = \begin{cases} 1, & y \geq \frac{\sigma^2}{2a} \ln\left(\frac{P[H_0]}{P[H_1]}\right) \\ 0, & y < \frac{\sigma^2}{2a} \ln\left(\frac{P[H_0]}{P[H_1]}\right) \end{cases} \quad \beta = \frac{\sigma^2}{2a} \ln\left(\frac{P[H_0]}{P[H_1]}\right)$$

$$P_e = P_{\text{FA}} P[H_0] + P_{\text{MD}} P[H_1]$$

$$P_{\text{FA}} = P[\{Y \in A_1\} | H_0] = P[Y \geq \beta | H_0] = \int_{\beta}^{\infty} f_{Y|H_0}(y) dy = Q\left(\frac{\beta - (-a)}{\sigma}\right) = Q\left(\frac{a + \beta}{\sigma}\right)$$

$$P_{\text{MD}} = P[\{Y \in A_0\} | H_1] = P[Y < \beta | H_1] = \int_{-\infty}^{\beta} f_{Y|H_1}(y) dy = \Phi\left(\frac{\beta - a}{\sigma}\right) = Q\left(\frac{a - \beta}{\sigma}\right)$$